

1 **Supplementary File 3: Additional Discussion**

2 **LEFT SHIFT IN CBWC BETWEEN V6.7.8.1 AND V8.0.6.2 GROW CALCULATORS**

3 We acknowledge that the median CBWC dropped (left shift) from 43% (v6.7.8.1
4 calculator) to 42% (v8.0.6.2 calculator) in the treatment group and from 31% (v6.7.8.1
5 calculator) to 28% (v8.0.6.2 calculator) in the sham group. The “Ghanaian” ethnicity
6 coefficient was not yet available in the v6.7.8.1 calculator, so the authors used the
7 regional “West African” coefficient for the original GPT analysis, as did we when we
8 replicated it (see Table 1 in our main text). The “West African” coefficient in the v6.7.8.1
9 calculator was derived from a 10 year database of routine, scan-dated, pregnancies
10 booked and delivered at a single unit (Queen’s Medical Centre, Nottingham, UK).(1)
11 While these data represent West Africans living in the UK, analyses have shown that
12 differences between 1st and 2nd generation migrants are negligible.(1) However, the
13 new centile calculators,(2) including the v8.0.6.2 calculator,(3) allow calculation of
14 CBWC and use an ultrasound-derived estimated fetal weight, for 120 ethnic, country of
15 origin, and region coefficients.(4) The v8.0.6.2 calculator is based on data from over 4
16 million pregnancies submitted from 33 countries.(4) The “Ghanaian” ethnicity coefficient
17 in the v8.0.6.2 calculator is, therefore, directly based on datasets from Ghana, not West
18 Africans living in the UK as with the v6.7.8.1 calculator. We submit that this key
19 difference between the old (v6.7.8.1) and new (v8.0.6.2) calculators likely accounts for
20 the left shift in the CBWC results. Furthermore, as GROW (Gestation-Related Optimal
21 Weight) sets an “optimal” standard, overall one would expect the CBWC distribution to
22 be left shifted and have more SGA than LGA cases, as it is more common for the
23 growth potential to be not reached than to be exceeded. For reasons cited above, we

24 also submit that our results with the new (v8.0.6.2) calculator are more representative of
25 the true population.

26 **COMPARISON OF FREQUENTIST AND BAYESIAN PARADIGMS**

27 **Null Findings and Sensitivity in the Frequentist Framework**

28 Note that in the course of experimentation, the null hypothesis is never proved or
29 established but is possibly disproved.(5) A type I statistical error, also known as a false
30 positive, occurs when investigators state that a treatment effect exists when it, in fact,
31 does not.(6) A type II statistical error, also known as a false negative, occurs when
32 investigators state that a treatment effect does not exist when, in fact, it does. When a
33 treatment is relatively safe, as is the case with positional therapy (PT), the
34 consequences of committing a type I error are minimal. In contrast, however, the
35 consequences of committing a type II error can be substantial because it may result in
36 patients being denied a treatment with beneficial effects, and may also suggest that
37 further research is not required.(7) Tragically, there are several examples of the latter
38 case occurring in the literature, including fibrinolytics.(8,9)

39 Whenever null findings are reported, it is important to complete a sensitivity analysis.
40 The 95% confidence interval (95% CI) of the sample mean difference, which was -0.3%
41 to 18.2% in our study, tells us the sensitivity of the experiment directly: if it includes both
42 the value of the null hypothesis (0%) and other interesting effect sizes, e.g., the
43 alternate hypothesis, then the experiment was not sensitive enough to draw definitive
44 conclusions.(10) This is inherent in the true meaning of the 95% CI, which is often
45 misinterpreted by readers of scientific literature as a de facto significance test by

46 examining whether its endpoints overlap the null value.(11) The 95% CI is computed by
47 finding the set of all values of the dependent variable (customised birthweight centile
48 [CBWC] in our study) that are non-significantly different from the sample mean
49 difference (8.4% in our study) at the 5% level. That is, all the points in our 95% CI (-
50 0.3% to 18.2%) are non-significantly different from the sample mean difference (8.4%),
51 which tells us that the GPT data are consistent with PT both having no effect (0%) and
52 with PT having an effect as large as an 18.2% gain in CBWC. Outside the interval of -
53 0.3% to 18.2%, out to infinity in both directions, all the values are significantly different
54 at the 5% level from the sample mean difference of 8.4% and can be ruled out as
55 possible population values. However, all the points within the 95% CI of the mean
56 difference in CBWC per the GROW standard (version 8.0.6.2) with PT versus sham-PT
57 (-0.3% to 18.2%) cannot be ruled out as population values. As discussed in
58 **Supplementary File 2**, while there is no consensus on the minimum clinically important
59 difference (MCID) in CBWC, our 95% CI contains effect sizes that most clinicians would
60 agree are important clinically such as 5%, 10%, and 15%. For example, a 5% gain in
61 CBWC would boost a growth restricted foetus (<3rd centile by definition) into the 3rd to
62 <10th centile range, halving its stillbirth risk, and a 10% gain in CBWC would boost it
63 into the normal range (10th to 90th centile), reducing its stillbirth risk by five fold.(12)

64 **Approach to Data and Hypotheses**

65 Approaching the data from a different analytical paradigm (Bayesian), however,
66 indicates that it is highly probably or almost certain that nightly maternal PT, compared
67 to sham-PT, during sleep throughout the third trimester confers a significant benefit to
68 foetal growth – even for the sceptic, benefit (>0% gain in CBWC) was highly probable

69 (0.81). This apparent discrepancy in the results under the frequentist paradigm
70 compared to the Bayesian paradigm is not so much a discrepancy in the results as it is
71 a discrepancy in the way that clinicians think and in the way these two paradigms
72 approach data and hypotheses.

73 The frequentist asks, “*Does my data fit my hypothesis?*”, whereas the Bayesian asks,
74 “*Does my hypothesis fit my data?*”. The p-value (frequentist paradigm) is the probability
75 that if the experiment were repeated an indefinite number of times, we would observe
76 results as extreme or more extreme than the results we observed assuming that the null
77 hypothesis is true (i.e., “*Does my data fit my hypothesis?*”).(10) As such, the p-value is a
78 long-run frequency; however, clinicians think in terms of conditional probabilities (i.e.,
79 “*Does my hypothesis fit my data?*”, or in other words, “*What is the probability of my*
80 *hypothesis being true given my data?*”), not long-run frequencies.(13)

81 When a clinician encounters a patient, they perform a history and physical exam (data)
82 and estimate a pre-test probability that the patient has a given disease (hypothesis).
83 The clinician orders further investigations of the patient, and based on these results
84 (new data), the post-test probability of disease (hypothesis) is revised, and so on. At no
85 point in this process is the clinician thinking about the patient, the data, and the
86 hypothesis in terms of long-run repeated experiments with a p-value.(6) Furthermore,
87 when it comes to treating a disease with an intervention, clinicians are more interested
88 in knowing the probability that the intervention is effective given the available data
89 (Bayesian thinking) and less interested in the long-run frequency of observing data as
90 extreme or more extreme than that observed previously assuming that the intervention
91 is ineffective (frequentist thinking).

92 **Justification of Reanalysis of a “Negative Trial” with Bayesian Methods**

93 A Bayesian reanalysis of a “negative” trial under the frequentist paradigm does not
94 create positive results nor should it be employed as a statistical alternative in an effort to
95 demonstrate a treatment effect. To demonstrate, it is important to note that if PT had no
96 effect, the posterior probability of PT being better than sham-PT in a Bayesian analysis
97 would be 0.50 on average. That is, a truly futile intervention (not helpful nor harmful)
98 would result in 50% of the area under the posterior probability curve (lighter shade) in
99 **Figure 2** (see **Results** section in main text) being to the left of the 0% threshold and
100 50% being to the right of the 0% threshold. This was not the case seen in our analysis
101 where the bulk of the area under the posterior probability curve – for a menu of priors
102 reflecting varying levels of enthusiasm and scepticism – is located to the right of the 0%
103 threshold, which indicates a treatment benefit. As such, we made a deliberate choice to
104 use Bayesian methods to reanalyze the GPT data because frequentist methods had
105 previously provided an incomplete summary of the results.

106 **LIMITATIONS**

107 **Exclusion of Birth Weight From Bayesian Analysis**

108 The original GPT publication had two primary outcomes: birth weight and CBWC.(14)
109 We did not complete a Bayesian analysis of the raw birth weight and, instead, chose the
110 CBWC for analysis for reasons described in the **Limitations** section of the main text.
111 That said, given the p-value (0.14), difference (gain of 110 grams), and 95% CI for the
112 treatment effect of PT on birth weight (–38 to 258 grams) in the original GPT

113 publication,(14) a Bayesian reanalysis of the GPT birth weights would likely show a
114 beneficial treatment effect of PT on birth weight as we saw with CBWC, especially in
115 light of the birth weight findings in Anderson et al.'s meta-analysis.(15) This meta-
116 analysis reported that for n=57 participants whose going-to-sleep position in the third
117 trimester (after 28 weeks) was supine, their infant's mean (standard deviation) birth
118 weight centile was 3410 (112) grams, and for n=1703 participants whose going-to-sleep
119 position in the third trimester was non-supine, their infant's mean (standard deviation)
120 birth weight centile was 3554 (98) grams. Comparing these two groups, the adjusted
121 mean difference in birth weight was a gain of 144 grams (95%CI: 36 to 253 grams; p-
122 value 0.009) with non-supine going-to-sleep position.

123 **Data-Derived Prior from Meta-analysis**

124 We used data from a recent individual participant data meta-analysis of sleeping
125 position in the third trimester (15) to inform our strongly enthusiastic (SE) prior (see the
126 **Methods** section in main text) for the Bayesian analysis. Following our methodology,
127 based on this meta-analysis, we would approach the GPT with a belief that nightly use
128 of PT from 28 weeks to birth increases CBWC by 9% on average (probability 0.50)
129 compared to sham-PT. This prior distribution corresponds to a probability of 0.99, 0.85,
130 and 0.40 that PT improves CBWC by 0%, 5%, and 10% or more compared to sham-PT.
131 This meta-analysis is a good source of information on which to base our prior belief vis-
132 a-vis the impact of sleeping position on CBWC for at least two reasons: first, it analyses
133 the CBWC using the customised growth standard that we used (GROW), and second,
134 its inclusion criteria stipulated that participants must be at 28 weeks gestation or greater

135 at the time of interview about their sleeping position, which was the same eligibility
136 criteria employed for participation in the GPT.

137 However, one major limitation is that the four studies included in the metaanalysis were
138 observational (case-control) studies and, as such, did not use PT and, rather, analysed
139 CBCW in light of the going-to-sleep position. While the going-to-sleep position does not
140 tell the whole story about what happens to sleeping position after sleep onset, a recent
141 study by Wilson et al. gives helpful and relevant insight: in the third trimester, the going-
142 to-sleep position is the dominant sleeping position overnight in the majority (54%) of
143 pregnancies, and the supine going-to-sleep position results in significantly more time
144 spent sleeping supine overnight compared to a non-supine going-to sleep position (48%
145 vs. 22.6%, $p < 0.001$).⁽¹⁶⁾ While these findings do not completely assuage this limitation
146 of informing our SE prior from Anderson et al.'s meta-analysis, it may attenuate it until
147 data are available from other interventional trials of PT in pregnancy. Note that as of the
148 writing of the main text of this manuscript, the GPT is the only interventional trial of PT
149 during sleep in pregnancy with foetal growth as an outcome.

150 Another limitation is that the participant samples of the four studies included in the
151 meta-analysis were drawn from populations in New Zealand, Australia, and the United
152 Kingdom, which are different from the participant sample in the GPT (Ghana). While the
153 GROW CBCW, which the meta-analysis used, accounts for maternal ethnicity and other
154 population-related factors (e.g., maternal height and weight), there may be other
155 population-related factors that affect foetal growth (e.g., socioeconomic status). As
156 such, sleeping position may not affect foetal growth in the same way in different
157 populations; however, one study from the same setting and centre as the GPT,

158 indicates the contrary, at least in the Ghanaian population, and found that the newborns
159 of participants who reported supine sleep during pregnancy were at a five-fold
160 increased risk of low birth weight.(17)

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